Regret Analysis of Stochastic and Nonstochastic Multi-armed Bandit Problems

Sébastien Bubeck Theory Group

Research



Part 1: i.i.d., adversarial, and Bayesian bandit models

Known parameters: number of arms n and (possibly) number of rounds $T \ge n$.

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Protocol: For each round $t=1,2,\ldots,T$, the player chooses $I_t \in [n]$ based on past observations and receives a reward/observation $Y_t \sim \nu_{I_t}$ (independently from the past).

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Performance measure: The cumulative regret is the difference between the player's accumulated reward and the maximum the player could have obtained had she known all the parameters,

$$\overline{R}_T = T\mu^* - \mathbb{E}\sum_{t\in[T]} Y_t.$$

Fundamental tension between **exploration** and **exploitation**. Many applications!



How small can we expect \overline{R}_T to be? Consider the 2-armed case where $\nu_1=\mathrm{Ber}(1/2)$ and $\nu_2=\mathit{Ber}(1/2+\xi\Delta)$ where $\xi\in\{-1,1\}$ is unknown.

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$$egin{array}{ll} \overline{R}_T(\xi=+1) + \overline{R}_T(\xi=-1) & \geq & \Delta au(T) + \Delta \sum_{t=1}^T \exp(- au(t)\Delta^2) \\ & \geq & \Delta \min_{t \in [T]} (t+T \exp(-t\Delta^2)) \\ & pprox & \frac{\log(T\Delta^2)}{\Delta}. \end{array}$$

See Bubeck, Perchet and Rigollet [2012] for the details.

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$$\overline{R}_T(\xi = +1) + \overline{R}_T(\xi = -1) \ge \Delta \tau(T) + \Delta \sum_{t=1}^T \exp(-\tau(t)\Delta^2)$$
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 $\approx \frac{\log(T\Delta^2)}{\Delta}.$

See Bubeck, Perchet and Rigollet [2012] for the details. For Δ fixed the lower bound is $\frac{\log(T)}{\Delta}$, and for the worse Δ ($\approx 1/\sqrt{T}$) it is \sqrt{T} (Auer, Cesa-Bianchi, Freund and Schapire [1995]: \sqrt{Tn} for the *n*-armed case).

Notation: $\Delta_i = \mu^* - \mu_i$ and $N_i(t)$ is the number of pulls of arm i up to time t. Then one has $\overline{R}_T = \sum_{i=1}^n \Delta_i \mathbb{E} N_i(T)$.

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Theorem (Lai and Robbins [1985])

Consider a strategy s.t. $\forall a > 0$, we have $\mathbb{E}N_i(T) = o(T^a)$ if $\Delta_i > 0$. Then for any Bernoulli distributions,

$$\liminf_{T\to +\infty} \frac{\overline{R}_T}{\log(T)} \geq \sum_{i:\Delta_i>0} \frac{\Delta_i}{\mathrm{kl}(\mu_i,\mu^*)}.$$

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Note that $\frac{1}{2\Delta_i} \geq \frac{\Delta_i}{\mathrm{kl}(\mu_i,\mu^*)} \geq \frac{\mu^*(1-\mu^*)}{2\Delta_i}$ so up to a variance-like term the Lai and Robbins lower bound is $\sum_{i:\Delta_i>0} \frac{\log(T)}{2\Delta_i}$.



Hoeffding's inequality: w.p. $\geq 1 - 1/T$, $\forall t \in [T], i \in [n]$,

$$\mu_i \leq \frac{1}{N_i(t)} \sum_{s < t: I_s = i} Y_s + \sqrt{\frac{2 \log(T)}{N_i(t)}} =: UCB_i(t).$$

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UCB (Upper Confidence Bound) strategy (Lai and Robbins [1985], Agarwal [1995], Auer, Cesa-Bianchi and Fischer [2002]):

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so that $\mathbb{E}N_i(T) \leq 2 + 8\log(T)/\Delta_i^2$ and in fact

$$\overline{R}_{\mathcal{T}} \leq 2 + \sum_{i:\Delta_i > 0} \frac{8\log(\mathcal{T})}{\Delta_i}.$$

i.i.d. multi-armed bandit: going further

1. Optimal constant (replacing 8 by 1/2 in the UCB regret bound) and Lai and Robbins variance-like term (replacing Δ_i by $\mathrm{kl}(\mu_i,\mu^*)$): see Cappé, Garivier, Maillard, Munos and Stoltz [2013].

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- 2. In many applications one is merely interested in *finding* the best arm (instead of maximizing cumulative reward): this is the best arm identification problem. For the fundamental strategies see Even-Dar, Mannor and Mansour [2006] for the fixed-confidence setting (see also Jamieson and Nowak [2014] for a recent short survey) and Audibert, Bubeck and Munos [2010] for the fixed budget setting. Key takeaway: one needs of order $\mathbf{H} := \sum_i \Delta_i^{-2}$ rounds to find the best arm.

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- 3. The UCB analysis extends to sub-Gaussian reward distributions. For heavy-tailed distributions, say with $1+\varepsilon$ moment for some $\varepsilon\in(0,1]$, one can get a regret that scales with $\Delta_i^{-1/\varepsilon}$ (instead of Δ_i^{-1}) by using a robust mean estimator, see Bubeck, Cesa-Bianchi and Lugosi [2012].

Adversarial multi-armed bandit, Auer, Cesa-Bianchi, Freund and Schapire [1995, 2001]

For $t=1,\ldots,T$, the player chooses $I_t\in[n]$ based on previous observations, and simultaneously an adversary chooses a loss vector $\ell_t\in[0,1]^n$. The player's loss/observation is $\ell_t(I_t)$.

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$$R_T = \max_{i \in [n]} \sum_{t \in [T]} (\ell_t(I_t) - \ell_t(i)), \quad \overline{R}_T = \max_{i \in [n]} \mathbb{E} \sum_{t \in [T]} (\ell_t(I_t) - \ell_t(i)).$$

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Obviously $\mathbb{E} R_T \geq \overline{R}_T$ and there is equality in the oblivious case (\equiv adversary's choice are independent of the player's choice). The case where ℓ_1,\ldots,ℓ_T is an i.i.d. sequence corresponds to the i.i.d. case we just studied. In particular we have a \sqrt{Tn} lower bound.

Exponential weights strategy for *full information* (ℓ_t is observed at the end of round t): play I_t at random from p_t where

$$p_{t+1}(i) = \frac{1}{Z_{t+1}} p_t(i) \exp(-\eta \ell_t(i)).$$

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$$\eta \sum_{t} \left(\sum_{j} p_{t}(i)\ell_{t}(i) - \ell_{t}(j) \right) = \operatorname{Ent}(\delta_{j} \| p_{1}) - \operatorname{Ent}(\delta_{j} \| p_{T+1}) + \sum_{t} \psi_{t}$$

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Exp3: replace ℓ_t by $\widetilde{\ell}_t$ in the exponential weights strategy, where

$$\widetilde{\ell}_t(i) = \frac{\ell_t(I_t)}{p_t(i)} \mathbb{1}\{i = I_t\}.$$

Key property: $\mathbb{E}_{I_t \sim p_t} \widetilde{\ell}_t(i) = \ell_t(i)$.

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Amazingly the variance term is automatically controlled:

$$\mathbb{E}_{I_t,I\sim p_t}\widetilde{\ell}_t(I)^2 \leq \mathbb{E}_{I_t,I\sim p_t}\frac{\mathbb{1}\{I=I_t\}}{p_t(I_t)^2} = \mathbb{E}_{I\sim p_t}\frac{1}{p_t(I)} = n.$$

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Thus with $\eta = \sqrt{2n\log(n)/T}$ one gets $\overline{R}_T \leq \sqrt{2Tn\log(n)}$.



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- 2. The extraneous logarithmic factor in the pseudo-regret upper can be removed, see Audibert and Bubeck [2009]. Conjecture: one cannot remove the log factor for the expected regret, that is for any strategy there exists an adaptive adversary such that $\mathbb{E}R_T = \Omega(\sqrt{Tn\log(n)})$.

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- 5. Graph feedback structure, regret with respect to *S* switches, label efficient, switching cost...



Set of models $\{(\nu_1(\theta), \dots, \nu_n(\theta)), \theta \in \Theta\}$ and prior distribution π_0 over Θ . The Bayesian regret is defined as

$$BR_T(\pi_0) = \mathbb{E}_{\theta \sim \pi_0} \overline{R}_T(\nu_1(\theta), \dots, \nu_n(\theta)).$$

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Notation: π_t denotes the posterior distribution on θ at time t.



Theorem (Gittins [1979])

Consider the product and γ -discounted case: $\Theta = \times_i \Theta_i$, $\nu_i(\theta) := \nu(\theta_i)$, $\pi_0 = \otimes_i \pi_0(i)$, and furthermore one is interested in maximizing $\mathbb{E} \sum_{t \geq 0} \gamma^t Y_t$.

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$$\sup \left\{ \lambda \in \mathbb{R} : \sup_{\tau} \mathbb{E} \left(\sum_{t < \tau} \gamma^t X_t + \frac{\gamma^\tau}{1 - \gamma} \lambda \right) \geq \frac{1}{1 - \gamma} \lambda \right\},$$

where the expectation is over (X_t) drawn from $\nu(\theta)$ with $\theta \sim \pi_s(i)$, and the supremum is taken over all stopping times τ .

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where the expectation is over (X_t) drawn from $\nu(\theta)$ with $\theta \sim \pi_s(i)$, and the supremum is taken over all stopping times τ . For much more (implementation for exponential families, interpretation as a multitoken Markov game, ...) see Dumitriu, Tetali and Winkler [2003], Gittins, Glazebrook, Weber [2011], Kaufmann [2014].

Weber [1992] gives an exquisite proof of Gittins theorem. Let

$$\lambda_t(i) := \sup \left\{ \lambda \in \mathbb{R} : \sup_{\tau} \mathbb{E} \sum_{t < \tau} \gamma^t (X_t - \lambda) \ge 0 \right\}$$

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- Since the prevailing charge is nonincreasing, the discounted sum of prevailing charge is maximized if we always pick the arm with maximum prevailing charge.
- 3. Gittins index does exactly 2. and that in this case 1. is an equality. Q.E.D.



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Theoretical guarantees for this highly practical strategy have long remained elusive. Recently Agrawal and Goyal [2012] and Kaufmann, Korda and Munos [2012] proved that TS with Bernoulli reward distributions and uniform prior on the parameters achieves $\overline{R}_T = O\left(\sum_i \frac{\log(T)}{\Delta_i}\right)$ (note that this is the frequentist regret!).

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Guha and Munagala [2014] conjecture that, for product priors, TS is a 2-approximation to the optimal Bayesian strategy for the objective of minimizing the number of pulls on suboptimal arms.

Assume a prior in the adversarial model, that is a prior over $(\ell_1, \ldots, \ell_T) \in [0, 1]^{n \times T}$, and let \mathbb{E}_t denote the posterior distribution (given $\ell_1(I_1), \ldots, \ell_{t-1}(I_{t-1})$).

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$$r_t(i) = \mathbb{E}_t(\ell_t(i) - \ell_t(i^*)), \text{ and } v_t(i) = \operatorname{Var}_t(\mathbb{E}_t(\ell_t(i)|i^*)).$$

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which implies:

$$\forall t, \mathbb{E}_{t} r_{t}(I_{t}) \leq \sqrt{C \ \mathbb{E}_{t} v_{t}(I_{t})}$$

$$\Rightarrow \ \mathbb{E} \sum_{t=1}^{T} r_{t}(I_{t}) \leq \sum_{t=1}^{T} \sqrt{C \ \mathbb{E} v_{t}(I_{t})}$$

$$\Rightarrow \ BR_{T} \leq \sqrt{C \ T \ H(i^{*})/2}.$$

Bayesian multi-armed bandit, accumulation of information

$$v_t(i) = \operatorname{Var}_t(\mathbb{E}_t(\ell_t(i)|i^*)), \ \pi_t(j) = \mathbb{P}_t(i^* = j), \ \mathbb{E}\sum_{t \in T} v_t(I_t) \leq \frac{1}{2}H(x^*)$$

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Equipped with Pinsker's inequality and basic information theory concepts (such as the mutual information \mathbb{I}) one has:

$$\begin{aligned} v_t(i) &= & \sum_{j} \pi_t(j) (\mathbb{E}_t(\ell_t(i)|i^* = j) - \mathbb{E}_t(\ell_t(i)))^2 \\ &\leq & \frac{1}{2} \sum_{j} \pi_t(j) \mathrm{Ent}(\mathcal{L}_t(\ell_t(i)|i^* = j) \|\mathcal{L}_t(\ell_t(i))) \\ &= & \frac{1}{2} \mathbb{I}_t(\ell_t(i), i^*) = H_t(i^*) - H_t(i^*|\ell_t(i)). \end{aligned}$$

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Thus $\mathbb{E}v_t(I_t) \leq \frac{1}{2}\mathbb{E}(H_t(i^*) - H_{t+1}(i^*)).$

Bayesian multi-armed bandit, TS' information ratio

Let
$$\bar{\ell}_t(i) = \mathbb{E}_t \ell_t(i)$$
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For TS the following shows that one can take C = n:

$$\mathbb{E}_{t}\bar{\ell}_{t}(I_{t}) - \sum_{i} \pi_{t}(i)\bar{\ell}_{t}(i,i) = \sum_{i} \pi_{t}(i)(\bar{\ell}_{t}(i) - \bar{\ell}_{t}(i,i))$$

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Thus TS always satisfies $BR_T \leq \sqrt{TnH(i^*)} \leq \sqrt{Tn\log(n)}$.



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Thus TS always satisfies $BR_T \leq \sqrt{TnH(i^*)} \leq \sqrt{Tn\log(n)}$. Side note: by the minimax theorem this implies there exists a strategy for the oblivious adversarial model with regret $\sqrt{Tn\log(n)}$.

Summary of basic results

- 1. In the i.i.d. model UCB attains a regret of $O\left(\sum_i \frac{\log(T)}{\Delta_i}\right)$ and by Lai and Robbins' lower bound this is optimal (up to a multiplicative variance term).
- 2. In the adversarial model Exp3 attains a regret of $O(\sqrt{Tn\log(n)})$ and this is optimal up to the logarithmic term.
- 3. In the Bayesian model, Gittins index gives an *optimal* strategy for the case of product priors. For general priors Thompson Sampling is a more flexible strategy. Its Bayesian regret is controlled by the entropy of the optimal decision. Moreover TS with an uninformative prior has frequentist guarantees comparable to UCB.

Part 2: Linear, non-linear, and contextual bandit

The linear bandit problem, Auer [2002]

Known parameters: compact action set $A \subset \mathbb{R}^n$, adversary's action set $\mathcal{L} \subset \mathbb{R}^n$, number of rounds \mathcal{T} .

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Protocol: For each round t = 1, 2, ..., T, the adversary chooses a loss vector $\ell_t \in \mathcal{L}$ and simultaneously the player chooses $a_t \in \mathcal{A}$ based on past observations and receives a loss/observation

$$Y_t = \ell_t^{\top} a_t.$$

$$R_T = \mathbb{E} \sum_{t=1}^T \ell_t^{\top} a_t - \min_{a \in \mathcal{A}} \mathbb{E} \sum_{t=1}^T \ell_t^{\top} a.$$

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Other models: In the i.i.d. model we assume that there is some underlying $\theta \in \mathcal{L}$ such that $\mathbb{E}(Y_t|a_t) = \theta^\top a_t$. In the Bayesian model we assume that we have a prior distribution ν over the sequence (ℓ_1,\ldots,ℓ_T) (in this case the expectation in R_T is also over $(\ell_1,\ldots,\ell_T) \sim \nu$). Alternatively we could assume a prior over θ .

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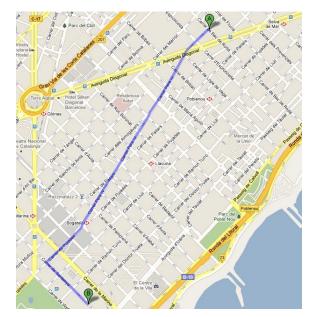
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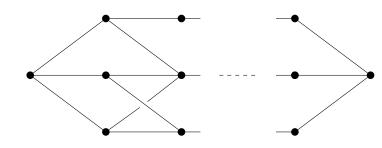
Assumption: unless specified otherwise we assume

$$\mathcal{L} = \mathcal{A}^{\circ} := \{\ell : \sup_{a \in \mathcal{A}} |\ell^{\top} a| \leq 1\}.$$



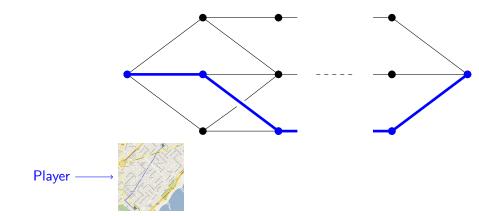


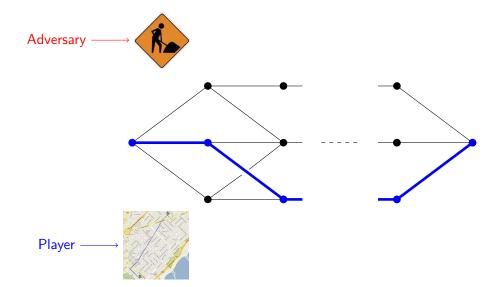
Adversary

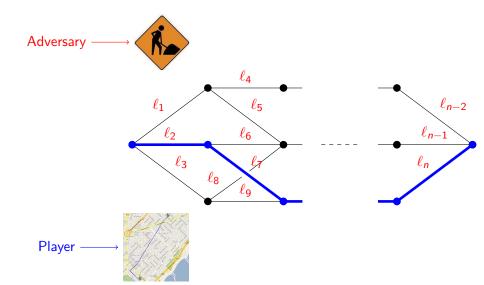


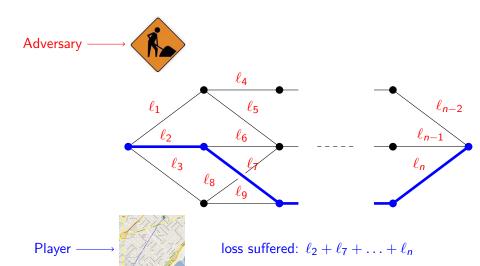
Player

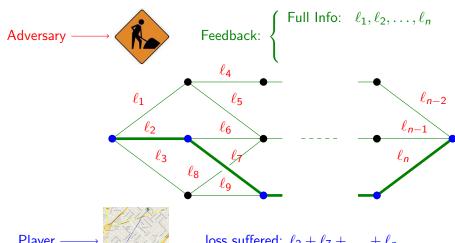
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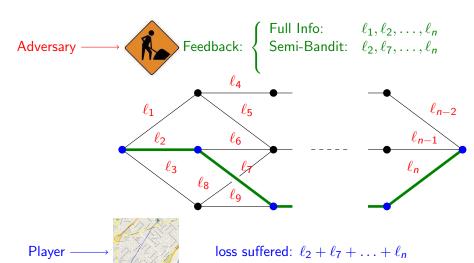


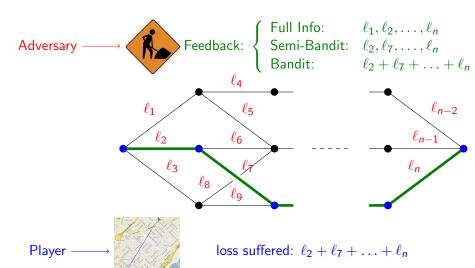






loss suffered: $\ell_2 + \ell_7 + \ldots + \ell_n$





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Using the eigenvalue formula for the trace and the Frobenius norm one can see that $Tr(M)^2 \le rank(M) ||M||_F^2$.

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Using the eigenvalue formula for the trace and the Frobenius norm one can see that $\mathrm{Tr}(M)^2 \leq \mathrm{rank}(M) \|M\|_F^2$. Moreover the rank of M is at most n since $M = UV^\top$ where $U, V \in \mathbb{R}^{|\mathcal{A}| \times n}$ (the i^{th} row of U is $\sqrt{\pi_t(i)}a_i$ and for V it is $\sqrt{\pi_t(i)}(\bar{\ell}_t - \bar{\ell}_t^i)$).

1. TS satisfies $R_T \leq \sqrt{nT \log(|\mathcal{A}|)}$. To appreciate the improvement recall that without the linear structure one would get a regret of order $\sqrt{|\mathcal{A}|T}$ and that \mathcal{A} can be exponential in the dimension n (think of the path planning example).

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- 2. Provided that one can efficiently sample from the posterior on ℓ_t (or on θ), TS just requires at each step one linear optimization over \mathcal{A} .
- 3. TS regret bound is optimal in the following sense. W.l.og. one can assume $|\mathcal{A}| \leq (10\,T)^n$ and thus TS satisfies $R_T = O(n\sqrt{T\log(T)})$ for any action set. Furthermore one can show that there exists an action set and a prior such that for any strategy one has $R_T = \Omega(n\sqrt{T})$, see Dani, Hayes and Kakade [2008], Rusmevichientong and Tsitsiklis [2010], and Audibert, Bubeck and Lugosi [2011, 2014].

Recall from Part 1 that exponential weights satisfies for any ℓ_t such that $\mathbb{E}\widetilde{\ell}_t(i) = \ell_t(i)$ and $\widetilde{\ell}_t(i) \geq 0$,

$$R_T \leq \frac{\max_i \operatorname{Ent}(\delta_i \| p_1)}{\eta} + \frac{\eta}{2} \mathbb{E} \sum_t \mathbb{E}_{I \sim p_t} \widetilde{\ell}_t(I)^2.$$

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DHK08 proposed the following (beautiful) unbiased estimator for the linear case:

$$\widetilde{\ell}_t = \Sigma_t^{-1} a_t a_t^\top \ell_t \text{ where } \Sigma_t = \mathbb{E}_{a \sim p_t} (a a^\top).$$

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Again, amazingly, the variance is automatically controlled:

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Up to the issue that $\widetilde{\ell}_t$ can take negative values this suggests the "optimal" $\sqrt{nT\log(|\mathcal{A}|)}$ regret bound.

1. The non-negativity issue of $\widetilde{\ell}_t$ is a manifestation of the need for an added exploration. DHK08 used a suboptimal exploration which led to an additional \sqrt{n} in the regret. This was later improved in Bubeck, Cesa-Bianchi, and Kakade [2012] with an exploration based on the John's ellipsoid (smallest ellipsoid containing \mathcal{A}).

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- 4. Bubeck and Eldan [2014]'s entropic barrier allows for a much more information-efficient sampling than AHR08. This gives another strategy with optimal regret which is efficient when \mathcal{A} is convex (and one can do linear optimization on \mathcal{A}).

Combinatorial setting: $\mathcal{A} \subset \{0,1\}^n$, $\max_a \|a\|_1 = m$, $\mathcal{L} = [0,1]^n$.

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- 4. Optimal regret in the semi-bandit case is \sqrt{mnT} and it can be achieved with mirror descent and the natural unbiased estimator for the semi-bandit situation.
- 5. For the bandit case the bound for exponential weights from the previous slides gives $m\sqrt{mnT}$. However the lower bound from ABL14 is $m\sqrt{nT}$, which is conjectured to be tight.

Assume $Y_t = \theta^\top a_t + \xi_t$ where (ξ_t) is an i.i.d. sequence of centered and sub-Gaussian real-valued random variables. The (regularized) least squares estimator for θ based on $\mathbb{Y}_t = (Y_1, \dots, Y_{t-1})^\top$ is, with $\mathbb{A}_t = (a_1 \dots a_{t-1}) \in \mathbb{R}^{n \times t-1}$ and $\Sigma_t = \lambda \mathrm{I}_n + \sum_{s=1}^{t-1} a_s a_s^\top$:

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A basic martingale argument (see e.g., Abbasi-Yadkori, Pál and Szepesvári [2011]) shows that w.p. $\geq 1 - \delta$, $\forall t \geq 1$,

$$\|\mathbb{A}_t \varepsilon_t\|_{\Sigma_t^{-1}} \leq \sqrt{\operatorname{logdet}(\Sigma_t) + \operatorname{log}(1/(\delta^2 \lambda^n))}.$$



Preliminaries for the i.i.d. case: a primer on least squares

Assume $Y_t = \theta^{\top} a_t + \xi_t$ where (ξ_t) is an i.i.d. sequence of centered and sub-Gaussian real-valued random variables. The (regularized) least squares estimator for θ based on $\mathbb{Y}_t = (Y_1, \dots, Y_{t-1})^{\top}$ is, with $\mathbb{A}_t = (a_1 \dots a_{t-1}) \in \mathbb{R}^{n \times t-1}$ and $\Sigma_t = \lambda I_n + \sum_{s=1}^{t-1} a_s a_s^{\top}$:

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Note that $\operatorname{logdet}(\Sigma_t) \leq n \operatorname{log}(\operatorname{Tr}(\Sigma_t)/n) \leq n \operatorname{log}(\lambda + t/n)$ (w.l.o.g. we assumed $||a_t|| \leq 1$).

Let $\beta = 2\sqrt{n\log(T)}$, and $\mathcal{E}_t = \{\theta' : \|\theta' - \hat{\theta}_t\|_{\Sigma_t} \leq \beta\}$. We showed that w.p. $\geq 1 - 1/T^2$ one has $\theta \in \mathcal{E}_t$ for all $t \in [T]$.

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The appropriate generalization of UCB is to select: $(\widetilde{\theta}_t, a_t) = \operatorname{argmin}_{(\theta', a) \in \mathcal{E}_t \times \mathcal{A}} \theta'^{\top} a$ (this optimization is NP-hard in general, more on that next slide).

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To control the sum of squares we observe that:

$$\det(\Sigma_{t+1}) = \det(\Sigma_t) \det(I_n + \Sigma_t^{-1/2} a_t (\Sigma_t^{-1/2} a_t)^\top) = \det(\Sigma_t) (1 + \|a_t\|_{\Sigma_t^{-1}}^2)$$
 so that (assuming $\lambda \ge 1$)

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Putting things together we see that the regret is $O(n \log(T) \sqrt{T})$.

So far we did not get any real benefit from the i.i.d. assumption (the regret guarantee we obtained is the same as for the adversarial model). To me the key benefit is in the simplicity of the i.i.d. algorithm which makes it easy to incorporate further assumptions.

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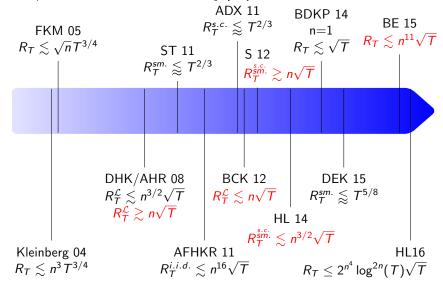
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- 4. $\log(T)$ -regime: if A is finite (note that a polytope is effectively finite for us) one can get $n^2 \log^2(T)/\Delta$ regret:

$$R_T \leq \mathbb{E} \sum_{t=1}^T \frac{(\theta^\top (a_t - a^*))^2}{\Delta} \leq \frac{\beta^2}{\Delta} \mathbb{E} \sum_{t=1}^T \|a_t\|_{\Sigma_t^{-1}}^2 \lesssim \frac{n^2 \log^2(T)}{\Delta}.$$

Some non-linear bandit problems

Lipschitz bandit: Kleinberg, Slivkins and Upfal [2008, 2016], Bubeck, Munos, Stoltz and Szepesvari [2008, 2011];

Gaussian process bandit: Srinivas, Krause, Kakade and Seeger [2010]; and convex bandit:



We now make the game-changing assumption that at the beginning of each round t a $context x_t \in \mathcal{X}$ is revealed to the player. The ideal notion of regret is now:

$$R_T^{\text{ctx}} = \sum_{t=1}^T \ell_t(a_t) - \inf_{\Phi: \mathcal{X} \to \mathcal{A}} \sum_{t=1}^T \ell_t(\Phi(x_t)).$$

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As far as I can tell the contextual bandit problem is an infinite playground and there is no canonical solution (or at least not yet!). Thankfully all we have learned so far can give useful guidance in this challenging problem.

Linear model after embedding

A natural assumption in several application domains is to suppose linearity in the loss after a correct embedding. Say we know mappings $(\varphi_a)_{a\in\mathcal{A}}$ such that $\mathbb{E}_t(\ell_t(a)) = \varphi_a(x_t)^\top \theta$ for some unknown $\theta \in \mathbb{R}^n$ (or in the adversarial case that $\ell_t(a) = \ell_t^\top \varphi_a(x_t)$).

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A much more challenging case is when the correct embedding $\varphi = (\varphi_a)_{a \in \mathcal{A}}$ is only known to belong to some class Φ . Without further assumptions on Φ we are basically back to the general model. Also note that a natural impulse is to run "bandits on top of bandits", that is first select some $\varphi_t \in \Phi$ and then select a_t based on the assumption that φ_t is correct. We won't get into this here, but let us investigate a related idea.

Exp4, Auer, Cesa-Bianchi, Freund and Schapire [2001]

One can play exponential weights on the set of policies with the following unbiased estimator (obvious notation: $\ell_t(\pi) = \ell_t(\pi(x_t))$, $\pi_t \sim p_t$, and $a_t = \pi_t(x_t)$)

$$\widetilde{\ell}_t(\pi) = \frac{\mathbb{1}\{\pi(x_t) = a_t\}}{\sum_{\pi': \pi'(x_t) = a_t} p_t(\pi')} \ell_t(a_t).$$

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Easy exercise: $R_T^{\rm ctx} \leq \sqrt{2T|\mathcal{A}|\log(|\Pi|)}$ (indeed the relative entropy term is smaller than $\log(|\Pi|)$ while the variance term is exactly $|\mathcal{A}|$).

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The only issue of this strategy is that the computationally complexity is linear in the policy space, which might be huge. A year and half ago a major paper by Agarwal, Hsu, Kale, Langford, Li and Schapire was posted, with a strategy obtaining the same regret as Exp4 (in the i.i.d. model) but which is also computationally efficient with an oracle for the offline problem (i.e., $\min_{\pi \in \Pi} \sum_{t=1}^{T} \ell_t(\pi(x_t))$). Unfortunately the algorithm is not simple enough yet to be included in these slides.

The statistician perspective, after Goldenshluger and Zeevi [2009, 2011], Perchet and Rigollet [2011]

Let $\mathcal{X} \subset \mathbb{R}^d$, $\mathcal{A} = [n]$, (x_t) i.i.d. from some μ absolutely continuous w.r.t. Lebesgue. The reward for playing arm a under context x is drawn from some distribution $\nu_a(x)$ on [0,1] with mean function $f_a(x)$ which is assumed to be β -Holder smooth. Let $\Delta(x)$ be the "gap" function.

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A key parameter is the proportion of contexts with a small gap. The margin assumption is that for some $\alpha>0$, one has

$$\mu(\lbrace x : \Delta(x) \in (0, \delta) \rbrace) \leq C\delta^{\alpha}, \forall \delta \in (0, 1].$$

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One can achieve a regret of order $T\left(\frac{n\log(n)}{T}\right)^{\frac{\beta(\alpha+1)}{2\beta+d}}$, which is optimal at least in the dependency on T. It can be achieved by running Successive Elimination on an adaptively refined partition of the space, see Perchet and Rigollet [2011] for the details.



The online multi-class classification perspective after Kakade, Shalev-Shwartz, and Tewari [2008]

Here the loss is assumed to be of the following very simple form: $\ell_t(a) = \mathbb{1}\{a \neq a_t^*\}$. In other words using the context x_t one has to predict the best action (which can be interpreted as a *class*) $a_t^* \in [n]$.

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KSST08 introduces the *banditron*, a bandit version of the multi-class perceptron for this problem. While with full information the online multi-class perceptron can be shown to satisfy a "regret" bound on of order \sqrt{T} , the banditron attains only a regret of order $T^{2/3}$. See also Chapter 4 in Bubeck and Cesa-Bianchi [2012] for more on this.

Summary of advanced results

- 1. The optimal regret for the linear bandit problem is $O(n\sqrt{T})$. In the Bayesian context Thompson Sampling achieves this bound. In the i.i.d. case one can use an algorithm based on the optimism in face of uncertainty together with concentration properties of the least squares estimator.
- 2. The i.i.d. algorithm can easily be modified to be computationally efficient, or to deal with sparsity in the unknown vector θ .
- 3. Extensions/variants: semi-bandit model, non-linear bandit (Lipschitz, Gaussian process, convex).
- 4. Contextual bandit is still a very active subfield of bandit theory.
- 5. Many important things were omitted. Example: knapsack bandit, see Badanidiyuru, Kleinberg and Slivkins [2013].



Some open problems we discussed

- 1. Prove the lower bound $\mathbb{E}R_T = \Omega(\sqrt{Tn\log(n)})$ for the adversarial *n*-armed bandit with adaptive adversary.
- Guha and Munagala [2014] conjecture: for product priors, TS is a 2-approximation to the optimal Bayesian strategy for the objective of minimizing the number of pulls on suboptimal arms.
- 3. Find a "simple" strategy achieving the Bubeck and Slivkins [2012] best of both worlds result.
- 4. For the combinatorial bandit problem, find a strategy with regret at most $n^{3/2}\sqrt{T}$ (current best is $n^2\sqrt{T}$).
- 5. Is there a computationally efficient strategy for i.i.d. linear bandit with optimal $n\sqrt{T}$ gap-free regret and with $\log(T)$ gap-based regret?
- 6. Is there a natural framework to think about "bandits on top of bandits" (while keeping \sqrt{T} -regret)?

